

A sequence of independent, identically distributed random variables (i.i.d.)

is a sequence $\{x_j\}_{j=1}^{\infty}$ s.t.

$\exists F : \mathbb{R} \rightarrow [0, 1]$, non-increasing, right-continuous and $\lim_{\lambda \rightarrow -\infty} F(\lambda) = 1$ and $\lim_{\lambda \rightarrow +\infty} F(\lambda) = 0$

s.t.

$$P(x_j > \lambda) = F(\lambda) \quad \text{for all } j \text{ and}$$

$$\begin{aligned} F_N(\lambda_1, \dots, \lambda_N) &= P(\{x_1 > \lambda_1, x_2 > \lambda_2, \dots, x_N > \lambda_N\}) \\ &= \prod_{j=1}^N P(x_j > \lambda_j) \\ &= \prod_{j=1}^N F(\lambda_j). \end{aligned}$$

An example of an i.i.d. is
the Rademacher sequence.

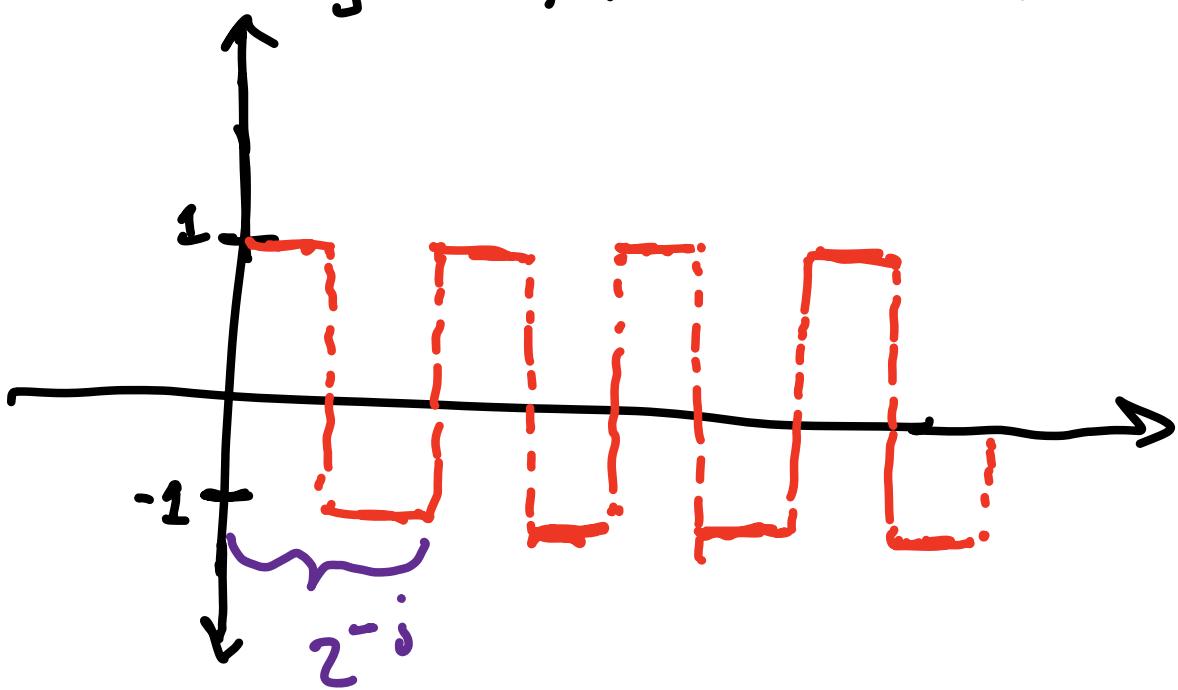
Rademacher Sequence

We encode a sequence of independent coin flips by the Rademacher functions on the measure space $([0,1], \mathcal{B}, \mathbb{P})$.

Define

$$r_j(t) := \text{sign}(\sin(2\pi 2^j t))$$

$$j = 0, 1, \dots, , \quad t \in [0, 1].$$



Lemma:

Let $\{r_j\}$ be the Rademacher sequence..

For $N \in \mathbb{Z}_+$ and $\{\alpha_j\}_{j=1}^N \subset \mathcal{A}$
one has

$$P\left(\left|\sum_{j=1}^N r_j \alpha_j\right| > \lambda \left(\sum_{j=1}^N |\alpha_j|^2\right)^{1/2}\right) \leq 4^{-\lambda^2/4}$$

For all $\lambda > 0$.

pf)

$$\begin{aligned} & P\left(\left|\sum_{j=1}^N r_j \alpha_j\right| > \lambda \left(\sum_{j=1}^N |\alpha_j|^2\right)^{1/2}\right) \\ & \leq P\left(\left|\sum_{j=1}^N r_j \operatorname{Re}(\alpha_j)\right| > \frac{\lambda}{2} \left(\sum_{j=1}^N |\alpha_j|^2\right)^{1/2}\right) \\ & \quad + P\left(\left|\sum_{j=1}^N r_j \operatorname{Im}(\alpha_j)\right| > \frac{\lambda}{2} \left(\sum_{j=1}^N |\alpha_j|^2\right)^{1/2}\right) \end{aligned}$$

For $t > 0$ and $S_N = \sum_{j=1}^N r_j \operatorname{Re}(\alpha_j)$

we have

$$\mathbb{E} e^{ts_N} = \prod_{j=1}^N \mathbb{E}(e^{tr_j Re a_j})$$

$$= \prod_{j=1}^N \cosh(t Re a_j)$$

Now note that $\cosh(x) \leq e^{x^2/2} \quad \forall x \in \mathbb{R}$

\Rightarrow

$$\mathbb{E} e^{ts_N} \leq \prod_{j=1}^N \exp\left(t^2 (Re a_j)^2 / 2\right)$$

$$\leq \exp\left(t^2 \frac{\sum |a_j|^2}{2}\right)$$

Then, if $\sigma = (\sum |a_j|^2)^{1/2}$

$$\mathbb{P}(s_N > \frac{\lambda}{2} \sigma)$$

$$= \mathbb{P}(e^{ts_N} \geq e^{\lambda t \sigma})$$

$$\leq \frac{\mathbb{E}[e^{ts_N}]}{e^{\lambda t \sigma}} \leq \frac{\exp\left(\frac{t^2}{2} \sum |a_j|^2\right)}{\exp(t \lambda \sigma / 2)}.$$

$$= \frac{\exp\left(\frac{t^2}{2} \sigma^2\right)}{\exp(t \lambda \sigma / 2)}$$

Let $t = \frac{\lambda}{2\sigma}$, then

$$\begin{aligned} P(\xi S_N > \frac{\lambda}{2}\sigma^2) &\leq \exp\left(\frac{\lambda^2}{8} - \frac{\lambda^2}{4}\right) \\ &= \exp(-\lambda^2/4) \end{aligned}$$

Similarly

$$P(\xi S_N < -\frac{\lambda}{2}\sigma^2) \leq \exp(-\lambda^2/4).$$

$$\Rightarrow P(|\xi S_N| > \frac{\lambda}{2}\sigma^2) \leq 2\exp(-\lambda^2/4)$$

Using the same argument for the imaginary part yields

$$P\{|\sum_{r;a_j}| > \lambda\sigma^2\} \leq 4\exp(-\lambda^2/4). \quad \square$$

The sub-Gaussian estimate implies all moments are equivalent